

CONTINUOUS DEPENDENCE OF THE SOLUTION OF RANDOM FRACTIONAL-ORDER DIFFERENTIAL EQUATION WITH NONLOCAL CONDITIONS

A. M. A. EL-SAYED, F. GAAFAR AND M. EL-GENDY

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Abstract. In this paper we are concerned with an initial value problem of random fractional-order differential equation with nonlocal condition. Continuous dependence and some other properties concerning the existence and uniqueness of the solution will be proved.

1. Introduction

The existence of solutions of nonlocal problems of differential equation of integer orders and fractional orders have been studied by some authors (see [4], [6], [7] and ([14]–[17]) for example).

Also stochastic problems are discussed in many papers, the reader is referred to ([1]-[3]), [5], ([8]-[13]) and [18].

Consider the random differential equation of fractional order

$$D^{\alpha}X(t) = c(t)f(X(t)) + b(t), \ t \in (0,T]$$
 (1)

with the nonlocal condition

$$X(0) + \sum_{k=1}^{n} a_k X(\tau_k) = X_0, \quad a_k > 0, \quad \tau_k \in (0, T),$$
 (2)

where X_0 is a second order random variable and a_k are positive real numbers.

Here we study the existence of a unique mean square continuous solution of the problem (1)–(2). The continuous dependence on the random variable X_0 and the deterministic coefficient a_k will be proved. The problem (1) with the integral condition

$$X(0) + \int_{0}^{T} X(s)dv(s) = X_{0}.$$
 (3)

will be considered. It must be noted that, in [4] the authors proved the existence of unique solution of the deterministic case of the problem (1)–(2).

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2. Preliminaries

Let I = [0,T] and $C = C(I,L_2(\Omega))$ be the class of all mean square continuous second order stochastic process with the norm

$$||X||_C = \sup_{t \in [0,T]} ||X(t)||_2 = \sup_{t \in [0,T]} \sqrt{E(X(t))^2}.$$

Now we have the following definitions.

DEFINITION 2.1. [4] Let $X \in C(I, L_2(\Omega))$ and $\beta \in \mathbb{R}^+$ The mean square fractional-order integral I_a^{β} of X(t) is defined by

$$I_a^{\beta}X(t) = \int_a^t \frac{(t-s)^{\beta-1}}{\Gamma(\beta)}X(s)ds$$

where $\Gamma(.)$ denotes the gamma function.

DEFINITION 2.2. [4] The mean square Caputo fractional-order derivative of order $\alpha \in (0,1]$ of the absolutely mean square continuous process X(t) is defined by

$$D_a^{\alpha}X(t) = I_a^{1-\alpha}\frac{d}{dt}X(t), \ 0 < \alpha \leqslant 1.$$

where D denotes the mean square differentiation and X(t) is assumed to be mean square differentiable.

3. Integral equation representation

Throughout the paper we assume that the following assumptions hold

(H1) the function f satisfies the mean square Lipschitz condition

$$|| f(X_1(t)) - f(X_2(t)) ||_2 \le k || X_1(t) - X_2(t) ||_2$$
.

(H2) There exists a positive real number m such that

$$\sup_{t\in[0,T]}|f(0)|\leqslant m.$$

(H3) c(t) and b(t) are absolutely continuous functions where

$$c = \sup_{t} |c(t)|, \quad b = \sup_{t} |b(t)|$$

Now we have the following lemma.

LEMMA 3.1. The solution of the nonlocal random problem (1) and (2) can be expressed by the integral equation

$$X(t) = a \left(X_0 - \sum_{k=1}^n a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} \right) + I^{\alpha} g(t, X(t))$$
 (4)

where
$$a = \left(1 + \sum_{k=1}^{n} a_{k}\right)^{-1}$$
.

Proof. For simplicity let

$$c(t)f(x(t)) + b(t) = g(t,x(t)).$$

If X(t) satisfies (1)–(2), then by using the properties of the stochastic fractional calculus, equation (1) can be written as

$$I^{1-\alpha}\dot{X}(t) = g(t, X(t)).$$

Operating by I^{α} on both sides in the last equation, we obtain

$$X(t) - X(0) = I^{\alpha}g(t, X(t))$$

substituting for the value of X(0) from equation (2), we obtain

$$X(t) = X_0 - \sum_{k=1}^{n} a_k X(\tau_k) + I^{\alpha} g(t, X(t)).$$

For $t = \tau_k$, we have

$$X(\tau_k) = X_0 - \sum_{k=1}^{n} a_k X(\tau_k) + I^{\alpha}, g(t, X(t))|_{t=\tau_k},$$

then

$$X(\tau_k) = X(t) - I^{\alpha}g(t, X(t)) + I^{\alpha}g(t, X(t))|_{t=\tau_k}.$$

So

$$\begin{split} X(t) &= X_0 - \sum_{k=1}^n a_k \left(X(t) - I^{\alpha} g(t, X(t)) + I^{\alpha} g(t, X(t)) |_{t=\tau_k} \right) + I^{\alpha} g(t, X(t)) \\ &= X_0 - \sum_{k=1}^n a_k X(t) + \sum_{k=1}^n a_k I^{\alpha} g(t, X(t)) - \sum_{k=1}^n a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} + I^{\alpha} g(t, X(t)) \end{split}$$

$$\left(1 + \sum_{k=1}^{n} a_k\right) X(t) = X_0 - \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t))|_{t=\tau_k} + \left(1 + \sum_{k=1}^{n} a_k\right) I^{\alpha} g(t, X(t)),$$

then

$$X(t) = \left(1 + \sum_{k=1}^{n} a_k\right)^{-1} \left(X_0 - \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t))|_{t=\tau_k}\right) + I^{\alpha} g(t, X(t))$$

and finally

$$X(t) = a\left(X_0 - \sum_{k=1}^n a_k I^{\alpha} g(t, X(t))|_{t=\tau_k}\right) + I^{\alpha} g(t, X(t))$$

where
$$a = \left(1 + \sum_{k=1}^{n} a_k\right)^{-1}$$
. \square

4. Existence and uniqueness

Now define the operator F by

$$FX(t) = a\left(X_0 - \sum_{k=1}^n a_k I^{\alpha} g(t, X(t))|_{t=\tau_k}\right) + I^{\alpha} g(t, X(t)).$$
(5)

Then we can prove the following lemma.

LEMMA 4.2. $F: C \rightarrow C$.

Proof. Let $X \in C$, t_1 , $t_2 \in [0,T]$ such that $|t_2 - t_1| < \delta$, then

$$FX(t_{2}) - FX(t_{1})$$

$$= \int_{0}^{t_{2}} \frac{(t_{2} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds - \int_{0}^{t_{1}} \frac{(t_{1} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds$$

$$= \int_{0}^{t_{1}} \frac{(t_{2} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds + \int_{t_{1}}^{t_{2}} \frac{(t_{2} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds$$

$$- \int_{0}^{t_{1}} \frac{(t_{1} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds$$

$$= \int_{0}^{t_{1}} \left[\frac{(t_{2} - s)^{\alpha - 1}}{\Gamma(\alpha)} - \frac{(t_{1} - s)^{\alpha - 1}}{\Gamma(\alpha)} \right] (c(s)f(X(s)) + b(s))ds$$

$$+ \int_{t_{1}}^{t_{2}} \frac{(t_{2} - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s))ds$$

and from assumption (H1) we have

$$|| f(X(t)) ||_2 - | f(0) | \le || f(X(t)) - f(0) ||_2 \le k || X(t) ||_2$$

then we have

$$|| f(X(t)) ||_2 \le k || X(t) ||_2 + | f(0) | \le k || X ||_C + m.$$

Then

$$\begin{split} &\|FX(t_2) - FX(t_1)\|_2 \\ & \leq \int_0^{t_1} \left[\frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} - \frac{(t_1 - s)^{\alpha - 1}}{\Gamma(\alpha)} \right] \mid c(s) \mid \|f(X(s))\|_2 \, ds \\ & + \int_0^{t_1} \left[\frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} - \frac{(t_1 - s)^{\alpha - 1}}{\Gamma(\alpha)} \right] \mid b(s) \mid ds \\ & + \int_{t_1}^{t_2} \frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \|f(X(s))\|_2 \, ds + \int_{t_1}^{t_2} \frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ & \leq c[k \parallel X \parallel_C + m] + b \int_0^{t_1} \left[\frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} - \frac{(t_1 - s)^{\alpha - 1}}{\Gamma(\alpha)} \right] ds \\ & + c[k \parallel X \parallel_C + m] + b \int_{t_1}^{t_2} \frac{(t_2 - s)^{\alpha - 1}}{\Gamma(\alpha)} ds \\ & = c[k \parallel X \parallel_C + m] + b) \left[-\frac{(t_2 - t_1)^{\alpha}}{\Gamma(\alpha + 1)} + \frac{t_2^{\alpha}}{\Gamma(\alpha + 1)} - \frac{t_1^{\alpha}}{\Gamma(\alpha + 1)} \right] \\ & + c[k \parallel X \parallel_C + m] + b) \left[\frac{(t_2 - t_1)^{\alpha - 1}}{\Gamma(\alpha)} \right] \\ & = c[k \parallel X \parallel_C + m] + b) \left[\frac{(t_2^{\alpha} - t_1^{\alpha})}{\Gamma(\alpha + 1)} \right] \\ & = \frac{c[k \parallel X \parallel_C + m] + b)}{\Gamma(\alpha + 1)} (t_2^{\alpha} - t_1^{\alpha}) \end{split}$$

which proves that $F: C \rightarrow C$. \square

For the existence of a unique continuous solution $X \in C$ of the problem (1)–(2), we have the following theorem.

THEOREM 4.1. Let the assumptions (H1)-(H3) be satisfied. If $A=\frac{kcT^{\alpha}}{\Gamma(\alpha+1)}<1$, then the problem (1)-(2) has a unique solution $X\in C$.

Proof. Let X and $Y \in C$, then

$$FX(t) - FY(t) = -a \sum_{k=1}^{n} a_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) (f(X(s)) - f(Y(s))) ds$$
$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) (f(X(s)) - f(Y(s))) ds$$

$$\begin{split} & \| FX(t) - FY(t) \|_2 \\ & \leqslant a \sum_{k=1}^n a_k \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) \| f(X(s)) - f(Y(s)) \|_2 \, ds \\ & + \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} | c(s) | \| f(X(s)) - f(Y(s)) \|_2 \, ds \\ & \leqslant ka \sum_{k=1}^n a_k \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} | c(s) | \| X(s) - Y(s) \|_2 \, ds \\ & + k \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} | c(s) | \| X(s) - Y(s) \|_2 \, ds \\ & \leqslant ka \| X - Y \|_C \sum_{k=1}^n a_k \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} | c(s) | \, ds + k \| X - Y \|_C \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} | c(s) | \, ds \\ & \leqslant ka \| X - Y \|_C \sum_{k=1}^n |a_k| \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \sup_s |c(s)| \, ds \\ & + k \| X - Y \|_C \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \sup_s |c(s)| \, ds \\ & \leqslant cka \| X - Y \|_C \sum_{k=1}^n a_k \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \, ds + ck \| X - Y \|_C \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \, ds \\ & \leqslant kc \| X - Y \|_C \left[a \sum_{k=1}^n a_k \int\limits_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \, ds + \int\limits_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \, ds \right] \\ & \leqslant kc \| X - Y \|_C \left[a \sum_{k=1}^n a_k \frac{\tau_k^{\alpha}}{\Gamma(\alpha + 1)} + \frac{t^{\alpha}}{\Gamma(\alpha + 1)} \right] \\ & \leqslant kc \| X - Y \|_C \left[a \sum_{k=1}^n a_k \frac{\tau_k^{\alpha}}{\Gamma(\alpha + 1)} + \frac{T^{\alpha}}{\Gamma(\alpha + 1)} \right] \end{aligned}$$

$$\leq \frac{kcT^{\alpha}}{\Gamma(\alpha+1)} \parallel X - Y \parallel_{C} \left[a \sum_{k=1}^{n} a_{k} + 1 \right] \leq \frac{2kcT^{\alpha}}{\Gamma(\alpha+1)} \parallel X - Y \parallel_{C},$$

then

$$\parallel FX - FY \parallel_C \leq 2A \parallel X - Y \parallel_C$$

where

$$A = \frac{kcT^{\alpha}}{\Gamma(\alpha+1)}.$$

Which proves that the operator F given by equation (5) is contraction. Consequently, the integral equation (4) has a unique fixed point $X \in C$.

To obtain the equivalence of equation (4) with the random initial value problem (1)–(2), let X be the solution of problem (1)–(2). As in [4] we can obtain

$$\frac{d}{dt}X(t) = \frac{d}{dt}I^{\alpha}g(t,X(t)) = \frac{d}{dt}\int_{0}^{t} \frac{(t-s)^{\alpha-1}}{\Gamma(\alpha)}g(s,X(s))ds$$
$$= g(t,X(t))|_{t=0}\frac{t^{\alpha-1}}{\Gamma(\alpha)} + I^{\alpha}\frac{d}{dt}g(t,X(t)).$$

Operate by $I^{1-\alpha}$ in both sides to obtain

$$D^{\alpha}X(t) = g(t,X(t)) = c(t)f(X(t)) + b(t).$$

For t = 0, we find the nonlocal condition (2) is satisfied as following, since we have

$$X(0) = a \left(X_0 - \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} \right)$$

and

$$X(\tau_k) = a\left(X_0 - \sum_{k=1}^n a_k I^{\alpha} g(t, X(t))|_{t=\tau_k}\right) + I^{\alpha} g(t, X(t))|_{t=\tau_k}$$

then

$$\begin{split} \sum_{k=1}^{n} a_k X(\tau_k) &= \sum_{k=1}^{n} a_k \left[a \left(X_0 - \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} \right) + I^{\alpha} g(t, X(t)) |_{t=\tau_k} \right] \\ &= a \sum_{k=1}^{n} a_k X_0 - a \sum_{k=1}^{n} a_k \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} + \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t)) |_{t=\tau_k} \end{split}$$

$$X(0) + \sum_{k=1}^{n} a_k X(\tau_k)$$

$$= a \left(X_0 - \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t))|_{t=\tau_k} \right) + a \sum_{k=1}^{n} a_k X_0$$

$$-a \sum_{k=1}^{n} a_k \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t))|_{t=\tau_k} + \sum_{k=1}^{n} a_k I^{\alpha} g(t, X(t))|_{t=\tau_k}$$

$$= X_0 \left(a + a \sum_{k=1}^n a_k \right) + \left(-a \sum_{k=1}^n a_k - a \sum_{k=1}^n a_k \sum_{k=1}^n a_k + \sum_{k=1}^n a_k \right) I^{\alpha} g(t, X(t))|_{t=\tau_k}$$

$$= aX_0 \left(1 + \sum_{k=1}^n a_k \right) + \left(-a \sum_{k=1}^n a_k \left(1 + \sum_{k=1}^n a_k \right) + \sum_{k=1}^n a_k \right) I^{\alpha} g(t, X(t))|_{t=\tau_k}$$

$$= aa^{-1}X_0 + \left(-aa^{-1} \sum_{k=1}^n a_k + \sum_{k=1}^n a_k \right) I^{\alpha} g(t, X(t))|_{t=\tau_k}$$

$$= X_0$$

which proves the equivalence. \Box

5. Continuous dependence

Consider the nonlocal random condition

$$X(0) + \sum_{k=1}^{n} a_k X(\tau_k) = \tilde{X}_0, \qquad \tau_k \in (0, T)$$
 (6)

DEFINITION 5.3. The solution $X \in C$ of the nonlocal random problem (1)–(2) is continuously dependent (on the data X_0) if $\forall \varepsilon > 0$, $\exists \delta > 0$ such that $\|X_0 - \tilde{X}_0\|_2 \le \delta$ implies that $\|X - \tilde{X}\|_C \le \varepsilon$.

Here, we study the continuous dependence (on the random data X_0) of the solution of the random fractional-order differential equation (1) and (2).

THEOREM 5.2. Let the assumptions (H1)-(H3) be satisfied. Then the solution of the nonlocal random problem (1)-(2) is continuously dependent on the random data X_0 .

Proof. Let

$$X(t) = a \left(X_0 - \sum_{k=1}^n a_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s)) ds \right)$$
$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s)) ds$$

be the solution of the nonlocal random problem (1)–(2) and

$$\tilde{X}(t) = a \left(\tilde{X}_0 - \sum_{k=1}^n a_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s) f(\tilde{X}(s)) + b(s)) ds \right)$$

$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s) f(\tilde{X}(s)) + b(s)) ds$$

be the solution of the nonlocal random problem (1) and (6). Then

$$X(t) - \tilde{X}(t)$$

$$= a(X_0 - \tilde{X}_0) - a \sum_{k=1}^n a_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) (f(X(s)) - f(\tilde{X}(s))) ds$$

$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) (f(X(s)) - f(\tilde{X}(s))) ds$$

Using our assumptions, we get

$$\begin{split} & \parallel X(t) - \tilde{X}(t) \parallel_{2} \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + a \sum_{k=1}^{n} a_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \parallel f(X(s)) - f(\tilde{X}(s)) \parallel_{2} ds \\ & + \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \parallel f(X(s)) - f(\tilde{X}(s)) \parallel_{2} ds \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + ka \sum_{k=1}^{n} a_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \parallel X(s) - \tilde{X}(s) \parallel_{2} ds \\ & + k \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \parallel X(s) - \tilde{X}(s) \parallel_{2} ds \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + ka \sum_{k=1}^{n} a_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} \sup_{s} \mid c(s) \mid \sup_{s} \parallel X(s) - \tilde{X}(s) \parallel_{2} ds \\ & + k \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \sup_{s} \mid c(s) \mid \sup_{s} \parallel X(s) - \tilde{X}(s) \parallel_{2} ds \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + kc \parallel X - \tilde{X} \parallel_{C} \left[a \sum_{k=1}^{n} a_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha + 1)} ds + \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} ds \right] \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + kc \parallel X - \tilde{X} \parallel_{C} \left[a \sum_{k=1}^{n} a_{k} \frac{T^{\alpha}}{\Gamma(\alpha + 1)} + \frac{T^{\alpha}}{\Gamma(\alpha + 1)} \right] \\ & \leqslant a \parallel X_{0} - \tilde{X}_{0} \parallel_{2} + kc \parallel X - \tilde{X} \parallel_{C} \frac{2T^{\alpha}}{\Gamma(\alpha + 1)}, \end{split}$$

then

$$\parallel X - \tilde{X} \parallel_C \leqslant \frac{A\delta}{(1 - 2A)}.$$

This complete the proof. \Box

Now consider the random fractional-order differential equation (1) with the nonlocal condition

$$X(0) + \sum_{k=1}^{n} \tilde{a}_k X(\tau_k) = X_0, \quad \tau_k \in (0, T)$$
 (7)

DEFINITION 5.4. The solution $X \in C$ of the nonlocal random problem (1)–(2) is continuously dependent (on the coefficient a_k of the nonlocal condition) if $\forall \varepsilon > 0$, $\exists \delta > 0$ such that $|a_k - \tilde{a}_k| \leq \delta$ implies that $||X - \tilde{X}||_C \leqslant \varepsilon$.

Here, we study the continuous dependence (on the coefficient a_k of the nonlocal condition) of the solution of the random fractional-order differential equation (1) and (2).

THEOREM 5.3. Let the assumptions (H1)-(H3) be satisfied. Then the solution of the nonlocal random problem (1)-(2) is continuously dependent on the coefficient a_k of the nonlocal condition.

Proof. Let

$$X(t) = a \left(X_0 - \sum_{k=1}^n a_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s)) ds \right)$$
$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(X(s)) + b(s)) ds$$

be the solution of the nonlocal random problem (1)–(2) and

$$\tilde{X}(t) = \tilde{a} \left(X_0 - \sum_{k=1}^n \tilde{a}_k \int_0^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(\tilde{X}(s)) + b(s)) ds \right)$$

$$+ \int_0^t \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} (c(s)f(\tilde{X}(s)) + b(s)) ds$$

be the solution of the nonlocal random problem (1) and (7).

Then

$$\begin{split} &X(t)-\tilde{X}(t)\\ &=(a-\tilde{a})X_0-a\sum_{k=1}^na_k\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds+\tilde{a}\sum_{k=1}^n\tilde{a}_k\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds\\ &-\left(a\sum_{k=1}^na_k-\tilde{a}\sum_{k=1}^n\tilde{a}_k\right)\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}b(s)ds+\int\limits_0^t\frac{(t-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)[f(X(s))-f(\tilde{X}(s))]ds \end{split}$$

Now we have

$$|a - \tilde{a}| = \left| \left(1 + \sum_{k=1}^{n} a_k \right)^{(-1)} - \left(1 + \sum_{k=1}^{n} \tilde{a}_k \right)^{(-1)} \right|$$

$$= \left| \sum_{k=1}^{n} (\tilde{a}_k - a_k) \left(1 + \sum_{k=1}^{n} a_k \right)^{(-1)} \left(1 + \sum_{k=1}^{n} \tilde{a}_k \right)^{(-1)} \right|$$

$$\leqslant \left| \sum_{k=1}^{n} (\tilde{a}_k - a_k) \right| \leqslant n\delta$$

and

$$\begin{split} &\tilde{a}\sum_{k=1}^{n}\tilde{a}_{k}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds-a\sum_{k=1}^{n}a_{k}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\ &=\tilde{a}\left(1+\sum_{k=1}^{n}\tilde{a}_{k}\right)\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds\\ &-a\left(1+\sum_{k=1}^{n}a_{k}\right)\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\ &-\tilde{a}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds+a\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\ &=\tilde{a}(\tilde{a}^{-1})\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds-a(a^{-1})\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\ &-\tilde{a}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds+a\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds \end{split}$$

then

$$\begin{split} &\tilde{a}\sum_{k=1}^{n}\tilde{a}_{k}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(\tilde{X}(s))ds-a\sum_{k=1}^{n}a_{k}\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\ &=-\int_{0}^{\tau_{k}}\frac{(\tau_{k}-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)[f(X(s))-f(\tilde{X}(s))]ds+a\int_{0}^{\tau_{k}}f(s,X(s))ds-\tilde{a}\int_{0}^{\tau_{k}}f(\tilde{X}(s))ds\\ &-\tilde{a}\int_{0}^{\tau_{k}}f(X(s))ds+\tilde{a}\int_{0}^{\tau_{k}}f(X(s))ds \end{split}$$

$$=-\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)[f(X(s))-f(\tilde{X}(s))]ds+[a-\tilde{a}]\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)f(X(s))ds\\+\tilde{a}\int\limits_0^{\tau_k}\frac{(\tau_k-s)^{\alpha-1}}{\Gamma(\alpha)}c(s)[f(X(s))-f(\tilde{X}(s))]ds.$$

From assumption (H2) we have

$$|| f(X(t)) ||_2 \le k || X(t) ||_2 + | f(0) | \le k || X ||_C + m.$$

also we have

$$\begin{split} & \left[a \sum_{k=1}^{n} a_k - \tilde{a} \sum_{k=1}^{n} \tilde{a}_k \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ &= \left[a \left(1 + \sum_{k=1}^{n} a_k \right) - \tilde{a} \left(1 + \sum_{k=1}^{n} \tilde{a}_k \right) \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ & - \left[a - \tilde{a} \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ &= \left[aa^{-1} - \tilde{a}\tilde{a}^{-1} \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds - \left[a - \tilde{a} \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ &= - \left[a - \tilde{a} \right] \int_{0}^{\tau_k} \frac{(\tau_k - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds, \end{split}$$

$$\| X(t) - \tilde{X}(t) \|_{2}$$

$$\leq |a - \tilde{a}| \| X_{0} \|_{2}$$

$$+ \left\| -a \sum_{k=1}^{n} a_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) f(X(s)) ds + \tilde{a} \sum_{k=1}^{n} \tilde{a}_{k} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) f(\tilde{X}(s)) ds \right\|_{2}$$

$$+ \left| a \sum_{k=1}^{n} a_{k} - \tilde{a} \sum_{k=1}^{n} \tilde{a}_{k} \right| \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} |b(s)| ds$$

$$+ \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} |c(s)| \| f(X(s)) - f(\tilde{X}(s)) \|_{2} ds$$

$$\begin{split} &\leqslant \mid a - \tilde{a} \mid \mid \mid X_{0} \mid \mid_{2} + \left| \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) [f(X(s)) - f(\tilde{X}(s))] ds \right| \Big|_{2} \\ &+ \left| \left| [a - \tilde{a}] \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) [f(X(s)) ds \right| \Big|_{2} \\ &+ \left| \tilde{a} \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} c(s) [f(X(s)) - f(\tilde{X}(s))] ds \right| \Big|_{2} \\ &+ \left| a \sum_{k=1}^{n} a_{k} - \tilde{a} \sum_{k=1}^{n} \tilde{a}_{k} \right| \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid b(s) \mid ds \\ &+ \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} \mid c(s) \mid \mid \mid f(X(s)) - f(\tilde{X}(s)) \mid |_{2} ds \\ &\leqslant n\delta \mid \mid X_{0} \mid \mid_{2} + kc \mid \mid X - \tilde{X} \mid \mid_{C} \int_{0}^{\tau_{k}} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} ds + cn\delta[k \mid \mid X \mid \mid_{C} + m] \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} ds \\ &+ c\tilde{a}k \mid \mid X - \tilde{X} \mid \mid_{C} \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} ds + nb\delta \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha)} ds \\ &+ kc \mid \mid X - \tilde{X} \mid \mid_{C} \int_{0}^{t} \frac{(t - s)^{\alpha - 1}}{\Gamma(\alpha)} ds \\ &\leqslant n\delta \left[\mid \mid X_{0} \mid \mid_{2} + c[k \mid \mid X \mid \mid_{C} + m] \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha + 1)} + b \int_{0}^{\tau_{k}} \frac{(\tau_{k} - s)^{\alpha - 1}}{\Gamma(\alpha + 1)} \right] + \frac{skcT^{\alpha}}{\Gamma(\alpha + 1)} \mid X - \tilde{X} \mid \mid_{C} \\ &\leqslant n\delta \left[\mid \mid X_{0} \mid \mid_{2} + c[k \mid \mid X \mid \mid_{C} + m] \frac{T^{\alpha}}{\Gamma(\alpha + 1)} + b \frac{T^{\alpha}}{\Gamma(\alpha + 1)} \right] + \frac{3kcT^{\alpha}}{\Gamma(\alpha + 1)} \mid X - \tilde{X} \mid \mid_{C} \end{aligned}$$

then

$$\|X - \tilde{X}\|_{C} \leqslant n\delta \left[\|X_0\|_2 + c[k \|X\|_C + m] \frac{T^{\alpha}}{\Gamma(\alpha + 1)} + b \frac{T^{\alpha}}{\Gamma(\alpha + 1)} \right] + 3A \|X - \tilde{X}\|_{C}$$

so

$$\parallel X - \tilde{X} \parallel_{C} \leq \frac{n\delta \left[\parallel X_{0} \parallel_{2} + c[k \parallel X \parallel_{C} + m] \frac{T^{\alpha}}{\Gamma(\alpha+1)} + b \frac{T^{\alpha}}{\Gamma(\alpha+1)} \right]}{(1 - 3A)}$$

This complete the proof.

6. Nonlocal integral condition

Let v(t) is a nondecreasing function such that

$$a_k = v(t_k) - v(t_{k-1}), \quad \tau_k \in (t_{k-1}, t_k)$$

where $0 < t_1 < t_2 < t_3 < \ldots < T$. Then, the nonlocal condition (2) will be in the form

$$X(0) + \sum_{k=1}^{n} X(\tau_k)(v(t_k) - v(t_{k-1})) = X_0.$$

From the mean square continuity of the solution of the nonlocal problem (1)–(2), we obtain from [18]

$$\lim_{n \to \infty} \sum_{k=1}^{n} X(\tau_k)(v(t_k) - v(t_{k-1})) = \int_{0}^{T} X(s) dv(s),$$

that is, the nonlocal conditions (2) is transformed to the mean square Riemann-Steltjes integral condition

$$X(0) + \int_{0}^{T} X(s)dv(s) = X_{0}.$$

Now, we have the following theorem.

THEOREM 6.4. Let the assumptions (H1)-(H3) be satisfied, then the stochastic differential equation (1) with the nonlocal integral condition (3) has a unique mean square continuous solution represented in the form

$$X(t) = a \left(X_0 - \int_0^T \int_0^s \frac{(\tau_k - \theta)^{\alpha - 1}}{\Gamma(\alpha)} (c(\theta) f(X(\theta)) + b(\theta)) d\theta dv(s) \right)$$
$$+ \int_0^t \frac{(t - \theta)^{\alpha - 1}}{\Gamma(\alpha)} (c(\theta) f(X(\theta)) + b(\theta)) d\theta$$

where $a^* = (1 + v(T) - v(0))^{-1}$.

Proof. Taking the limit of equation (3) we get the proof. \Box

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A. M. A. El-Sayed Faculty of Science Alexandria University Egypt

e-mail: amasyed@alexu.edu.eg

F. Gaafar Faculty of Science Damanhour University Egypt

 $e ext{-}mail:$ fatmagaafar2@yahoo.com

M. El-Gendy Faculty of Science Damanhour University Egypt

e-mail: maysa_elgendy@yahoo.com