

## AN EXTENSION OF THE DAVIS–GUT LAW AND LAI LAW

YONG ZHANG

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*Abstract.* Let  $\{X, X_n, n \geq 1\}$  be a sequence of i.i.d. random variables with  $EX = 0$  and  $EX^2 = 1$  and the partial sums  $S_n = \sum_{k=1}^n X_k$ ,  $n \geq 1$ . Assume that  $f(x)$  and  $g(x)$  are positive functions defined on  $[0, \infty)$ . In this short note, under some suitable conditions, we establish the following results

$$\sum_{n=1}^{\infty} f(n)P\{|S_n| > \beta\sqrt{ng(n)}\} < \infty \text{ or } = \infty$$

according as

$$\sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2}{2}g^2(n)(1 + \alpha(n))\right\} < \infty \text{ or } = \infty$$

where  $\alpha(n) = EX^2I\{|X| > \sqrt{ng(n)}\}/EX^2I\{|X| \leq \sqrt{ng(n)}\}$ ,  $\beta > 0$ . The results extend and generalize the known Davis-Gut Law and Lai Law.

### 1. Introduction and main results

Throughout this paper, let  $\{X, X_n, n \geq 1\}$  be a sequence of i.i.d. random variables with  $EX = 0$ ,  $EX^2 = 1$  and denote  $S_n = \sum_{k=1}^n X_k$ ,  $n \geq 1$ . Let  $\mathcal{N}$  be the standard normal random variable.  $C$  denotes a positive constant, possibly varying from place to place, the notion  $a_n \sim b_n$  stands for  $\lim_{n \rightarrow \infty} a_n/b_n = 1$ .

The following theorem, which was related to the classical Hartman-Wintner law of the iterated logarithm, is labeled as the Davis-Gut Law.

**THEOREM A.** *Let  $\{X, X_n, n \geq 1\}$  be a sequence of i. i. d. random variables with the partial sums  $S_n = \sum_{k=1}^n X_k$ ,  $n \geq 1$ . Then the following three statements are equivalent*

$$EX = 0 \text{ and } EX^2 = 1, \tag{1.1}$$

$$\sum_{n=1}^{\infty} \frac{1}{n} P\{|S_n| > (1 + \varepsilon)\sqrt{2n \log \log n}\} \begin{cases} < \infty, \text{ if } \varepsilon > 0, \\ = \infty, \text{ if } \varepsilon < 0. \end{cases} \tag{1.2}$$

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$$\sum_{n=1}^{\infty} \frac{\log \log n}{n} P\{|S_n| > (1 + \varepsilon)\sqrt{2n \log \log n}\} \begin{cases} < \infty, \text{ if } \varepsilon > 0, \\ = \infty, \text{ if } \varepsilon < 0. \end{cases} \tag{1.3}$$

(1.1)  $\Rightarrow$  (1.2) can be deduced by Theorem 4 of Davis [3] which was remedied by Corollary 2.3 of Li et al. [7], (1.2)  $\Rightarrow$  (1.1) can be got by Theorem 6.2 of Gut [4]. (1.1)  $\Leftrightarrow$  (1.3) was proved by Li [6]. For analogue results, we refer to Li and Rosalsky [8], Liu et al. [9].

The next theorem, which was related to the law of the single logarithm, is labeled as the Lai Law.

**THEOREM B.** *Let  $\{X, X_n, n \geq 1\}$  be a sequence of i. i. d. random variables with the partial sums  $S_n = \sum_{k=1}^n X_k, n \geq 1$ . Then the following two statements are equivalent*

$$EX = 0, EX^2 = 1 \text{ and } E(X^2/\log|X|)^{r+1} < \infty, r > 0, \tag{1.4}$$

$$\sum_{n=1}^{\infty} n^{r-1} P\{|S_n| > (1 + \varepsilon)\sqrt{2rn \log n}\} \begin{cases} < \infty, \text{ if } \varepsilon > 0, \\ = \infty, \text{ if } \varepsilon < 0. \end{cases} \tag{1.5}$$

Lai [5] established the equivalence for  $\varepsilon > 0$ , Chen and Wang [2] extended it to  $\varepsilon < 0$ . Chen and Qi [1] discussed the similar results of (1.5) for  $r = 0$ . Liu and Meng [11] established the analogue results of Davis-Gut Law and Lai Law for finitely inhomogeneous walks.

The gap for  $\varepsilon = 0$  of Theorem A was solved by Liu and Guo [10], they obtained

**THEOREM C.** *Let  $\{X, X_n, n \geq 1\}$  be a sequence of i. i. d. random variables with  $EX = 0, EX^2 = 1$  and the partial sums  $S_n = \sum_{k=1}^n X_k, n \geq 1$ . Then*

$$\sum_{n=1}^{\infty} \frac{1}{n} P\{|S_n| > \sqrt{2n \log \log n}\} < \infty \text{ or } = \infty \tag{1.6}$$

according as

$$\sum_{n=1}^{\infty} \frac{1}{n\sqrt{\log \log n}} \cdot \frac{1}{(\log n)^{1+\alpha(n)}} < \infty \text{ or } = \infty \tag{1.7}$$

where  $\alpha(n) = EX^2 I\{|X| > \sqrt{n \log \log n}\} / EX^2 I\{|X| \leq \sqrt{n \log \log n}\}$ .

$$\sum_{n=1}^{\infty} \frac{\log \log n}{n} P\{|S_n| > \sqrt{2n \log \log n}\} < \infty \text{ or } = \infty \tag{1.8}$$

according as

$$\sum_{n=1}^{\infty} \frac{\sqrt{\log \log n}}{n(\log n)^{1+\alpha(n)}} < \infty \text{ or } = \infty \tag{1.9}$$

where  $\alpha(n) = EX^2 I\{|X| > \sqrt{n \log \log n}\} / EX^2 I\{|X| \leq \sqrt{n \log \log n}\}$ .

In this paper, we want to extend and generalize the above Theorems to some functions. The main result of this note is the following.

**THEOREM 1. (Main)** *Let  $\{X, X_n, n \geq 1\}$  be a sequence of i. i. d. random variables with  $EX = 0$ ,  $EX^2 = 1$  and the partial sums  $S_n = \sum_{k=1}^n X_k$ ,  $n \geq 1$ . Let  $f(x)$ ,  $g(x)$  be positive functions defined on  $[0, \infty)$ , suppose that the following conditions hold*

(C1)  $g(x)$  is strictly increasing to  $\infty$ ,

(C2)  $\sum_{k=1}^n kf(k) \leq Cn^2f(n)$ ,

(C3)  $\sum_{k=n}^{\infty} k^{-1/2}f(k)g^{-3}(k) \leq Cn^{1/2}f(n)g^{-3}(n)$ ,

(C4)  $nf(n) = O(g^2(n))$ , or

(C4\*)  $\limsup_{n \rightarrow \infty} \frac{nf(n)}{g^2(n)} = \infty$ ,  $\limsup_{n \rightarrow \infty} \max_{ng^2(n) \leq x^2 \leq (n+1)g^2(n+1)} \frac{g^4(x^2)}{g^4(n)} \cdot \frac{f(n)}{f(\frac{x^2}{g^2(n)})} \leq C$ ,

$E[\frac{X^4}{g^4(X^2)}f(\frac{X^2}{g^2(X^2)})] < \infty$ . Then

$$\sum_{n=1}^{\infty} f(n)P\{|S_n| > \beta\sqrt{ng(n)}\} < \infty \text{ or } = \infty \tag{1.10}$$

according as

$$\sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\{-\frac{\beta^2}{2}g^2(n)(1 + \alpha(n))\} < \infty \text{ or } = \infty \tag{1.11}$$

where  $\alpha(n) = EX^2I\{|X| > \sqrt{ng(n)}\}/EX^2I\{|X| \leq \sqrt{ng(n)}\}$ ,  $\beta > 0$ .

**REMARK 1.** There are many functions satisfying the assumptions of  $f(x)$  and  $g(x)$ , such as  $f(x) = \frac{l(x)}{x^\alpha}$ ,  $g(x) = x^\gamma h(x)$ , with some suitable conditions of  $\alpha$ ,  $\gamma \geq 0$  and  $l(x)$  and  $h(x)$  are slowly varying at infinity. We list some examples in the following Corollaries.

**COROLLARY 1.** *Let  $f(n) = \frac{(\log \log n)^b}{n}$ ,  $g(n) = \sqrt{\log \log n}$  with  $EX^2(\log \log |X|)^{\max\{0, b-1\}} < \infty$ ,  $b \in R$ , then*

$$\sum_{n=1}^{\infty} \frac{(\log \log n)^b}{n} P\{|S_n| > \beta\sqrt{n \log \log n}\} \begin{cases} < \infty, \text{ if } \beta > \sqrt{2}, \\ = \infty, \text{ if } \beta < \sqrt{2}. \end{cases}$$

And for  $\beta = \sqrt{2}$ , we have

$$\sum_{n=1}^{\infty} \frac{(\log \log n)^b}{n} P\{|S_n| > \sqrt{2n \log \log n}\} < \infty \text{ or } = \infty \tag{1.12}$$

according as

$$\sum_{n=1}^{\infty} \frac{(\log \log n)^{b-1/2}}{n(\log n)^{1+\alpha(n)}} < \infty \text{ or } = \infty \tag{1.13}$$

where  $\alpha(n) = EX^2I\{|X| > \sqrt{n \log \log n}\}/EX^2I\{|X| \leq \sqrt{n \log \log n}\}$ .

COROLLARY 2. Let  $f(n) = \frac{(\log n)^b}{n}$ ,  $g(n) = \sqrt{\log \log n}$  with  $EX^2 = 1$  for  $b \leq 0$  and  $EX^2(\log |X|)^b(\log \log |X|)^{-1} < \infty$  for  $b > 0$ , then

$$\sum_{n=1}^{\infty} \frac{(\log n)^b}{n} P\{|S_n| > \beta \sqrt{n \log \log n}\} \begin{cases} < \infty, & \text{if } \beta^2 > 2(1+b), \\ = \infty, & \text{if } \beta^2 < 2(1+b). \end{cases}$$

And for  $\beta^2 = 2(1+b) > 0$ , we have

$$\sum_{n=1}^{\infty} \frac{(\log n)^b}{n} P\{|S_n| > \sqrt{2(1+b)n \log \log n}\} < \infty \text{ or } = \infty \quad (1.14)$$

according as

$$\sum_{n=1}^{\infty} \frac{1}{n(\log \log n)^{1/2}} \frac{1}{(\log n)^{1+(1+b)\alpha(n)}} < \infty \text{ or } = \infty \quad (1.15)$$

where  $\alpha(n) = EX^2 I\{|X| > \sqrt{n \log \log n}\} / EX^2 I\{|X| \leq \sqrt{n \log \log n}\}$ .

COROLLARY 3. Let  $f(n) = \frac{(\log \log n)^b}{n}$ ,  $g(n) = \log^s n$  with  $EX^2 = 1$ ,  $s > 0$ , then we have

$$\sum_{n=1}^{\infty} \frac{(\log \log n)^b}{n} P\{|S_n| > \beta \sqrt{n} \log^s n\} < \infty.$$

COROLLARY 4. Let  $f(n) = n^{r-1}$ ,  $g(n) = \sqrt{\log n}$  with  $E[\frac{X^2}{\log |X|}]^{r+1} < \infty$ ,  $0 < r < 1/2$ , then

$$\sum_{n=1}^{\infty} n^{r-1} P\{|S_n| > \beta \sqrt{n \log n}\} \begin{cases} < \infty, & \text{if } \beta > \sqrt{2r}, \\ = \infty, & \text{if } \beta < \sqrt{2r}. \end{cases}$$

And for  $\beta = \sqrt{2r}$ , we have

$$\sum_{n=1}^{\infty} n^{r-1} P\{|S_n| > \sqrt{2rn \log n}\} < \infty \text{ or } = \infty \quad (1.16)$$

according as

$$\sum_{n=1}^{\infty} \frac{1}{n^{1+r\alpha(n)}(\log n)^{1/2}} < \infty \text{ or } = \infty. \quad (1.17)$$

where  $\alpha(n) = EX^2 I\{|X| > \sqrt{n \log n}\} / EX^2 I\{|X| \leq \sqrt{n \log n}\}$ .

COROLLARY 5. Let  $f(n) = n^{-\alpha}$ ,  $g(n) = n^\gamma$ ,  $0 < \alpha \leq 1$ ,  $\gamma > 0$ ,  $\alpha + 2\gamma \geq 1$ ,  $EX^2 = 1$ , then

$$\sum_{n=1}^{\infty} \frac{1}{n^\alpha} P\{|S_n| > \beta \sqrt{nn}^\gamma\} < \infty.$$

Especially, for  $0 < p < r < 2$ ,

$$\sum_{n=1}^{\infty} n^{\frac{r}{p}-2} P\{|S_n| > \beta n^{\frac{1}{p}}\} < \infty.$$

REMARK 2. Obviously, Corollaries 1-5 extend the known results (Theorems A-C), and we also give the gap  $\varepsilon = 0$  in Theorem B.

The following examples shows that it is possible that the series converge or diverge in (1.13), (1.15) and (1.17) under some different conditions.

EXAMPLE 1. Let  $X$  be a random variable such that  $\sqrt{c/d}X$  has the following density function

$$f_z(x) = \frac{1}{c} \cdot \frac{\log \log \log |x|}{|x|^3 \log |x| (\log \log |x|)^2} I\{|x| > z\}, \quad z > 0$$

where

$$c = c(z) = 2 \int_z^{\infty} \frac{\log \log \log x}{x^3 \log x (\log \log x)^2} dx, \quad d = d(z) = 2 \int_z^{\infty} \frac{\log \log \log x}{x \log x (\log \log x)^2} dx.$$

By the same discussion in Liu and Guo [10], we know  $d/c$  is strictly monotone increasing in  $(0, \infty)$  and  $\lim_{z \rightarrow 0} d/c = 0$  and  $\lim_{z \rightarrow \infty} d/c = \infty$ , and

$$EX = 0, \quad EX^2 = 1, \quad EX^2 I\{|X| > \sqrt{n \log \log n}\} \sim \frac{d}{c} \cdot \frac{\log \log \log n}{\log \log n},$$

$$(\log n)^{EX^2 I\{|X| > \sqrt{n \log \log n}\}} \sim (\log \log n)^{\frac{d}{c}}, \quad n \rightarrow \infty.$$

Therefore there exist unique constants  $z_1$  and  $z_2$  such that

$$\frac{d(z_1)}{c(z_1)} = \frac{1}{2} + b > 0, \quad \frac{d(z_2)}{c(z_2)} = \frac{1}{2(1+b)} > 0.$$

So we can easily conclude the series (1.13) converges or diverges according as  $z \in (z_1, \infty)$  or  $z \in (0, z_1)$ , the series (1.15) converges or diverges according as  $z \in (z_2, \infty)$  or  $z \in (0, z_2)$ .

EXAMPLE 2. As the same argument as in Example 1,  $\sqrt{c'/d'}Y$  has the following density function

$$f_z(y) = \frac{1}{c'} \cdot \frac{\log \log |y|}{|y|^3 (\log |y|)^2} I\{|y| > z\}, \quad z > 0$$

where

$$c' = c'(z) = 2 \int_z^{\infty} \frac{\log \log x}{x^3 (\log x)^2} dx, \quad d' = d'(z) = 2 \int_z^{\infty} \frac{\log \log x}{x (\log x)^2} dx.$$

It is easy to prove  $d'/c'$  is strictly monotone increasing in  $(0, \infty)$  and  $\lim_{z \rightarrow 0} d'/c' = 0$  and  $\lim_{z \rightarrow \infty} d'/c' = \infty$ , and

$$EY = 0, \quad EY^2 = 1, \quad EY^2 I\{|Y| > \sqrt{n \log n}\} \sim \frac{d'}{c'} \cdot \frac{\log \log n}{\log n},$$

$$n^{EX^2} I\{|X| > \sqrt{n \log n}\} \sim (\log n)^{\frac{d'}{c'}}, \quad n \rightarrow \infty.$$

Therefore there exists an unique constant  $z_3$  such that

$$\frac{d'(z_3)}{c'(z_3)} = \frac{1}{2r}.$$

So we can easily conclude the series (1.17) converges or diverges according as  $z \in (z_3, \infty)$  or  $z \in (0, z_3)$ .

### 2. Proof of Theorem 1

The following lemmas are useful for the proof of Theorem 1. The first lemma is the nonuniform estimates of the remainder term in the central limit theorem (see Nagaev [12], Theorem 3, p.215).

LEMMA 1. *Let  $\{X, X_n, n \geq 1\}$  be a sequence of i.i.d. random variables with  $EX = 0, EX^2 = \sigma^2$  and  $E|X|^q < \infty, 2 < q \leq 3$ , then there a positive constant  $C$ , such that*

$$\left| P\left\{ \frac{S_n}{\sigma\sqrt{n}} \leq x \right\} - P\{\mathcal{N} \leq x\} \right| \leq \frac{CE|X|^q}{n^{q/2-1}(1+|x|^q)}.$$

LEMMA 2. *Under the conditions of Theorem 1, we have*

$$\sum_{n=1}^{\infty} nf(n)P\{|X| > \sqrt{ng(n)}\} < \infty, \tag{2.1}$$

$$\sum_{n=1}^{\infty} \frac{f(n)}{\sqrt{ng^3(n)}} E|X|^3 I\{|X| \leq \sqrt{ng(n)}\} < \infty. \tag{2.2}$$

*Proof.* Obviously, note that (C2),

$$\begin{aligned} & \sum_{n=1}^{\infty} nf(n)P\{|X| > \sqrt{ng(n)}\} = \sum_{n=1}^{\infty} nf(n) \sum_{k=n}^{\infty} P\{\sqrt{kg}(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &= \sum_{k=1}^{\infty} P\{\sqrt{kg}(k) < |X| \leq \sqrt{k+1}g(k+1)\} \sum_{n=n_0}^k nf(n) \\ &\leq C \sum_{k=1}^{\infty} k^2 f(k) P\{\sqrt{kg}(k) < |X| \leq \sqrt{k+1}g(k+1)\}. \end{aligned}$$

If (C4) holds, we have

$$\begin{aligned} \sum_{n=1}^{\infty} nf(n)P\{|X| > \sqrt{ng(n)}\} &\leq C \sum_{k=1}^{\infty} kg^2(k)P\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq CE|X|^2 < \infty. \end{aligned}$$

If (C4\*) holds, we have

$$\begin{aligned} \sum_{n=1}^{\infty} nf(n)P\{|X| > \sqrt{ng(n)}\} &\leq C \sum_{k=1}^{\infty} k^2 f(k)P\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &= C \sum_{k=1}^{\infty} E\left[\frac{X^4}{g^4(X^2)} f\left(\frac{X^2}{g^2(X^2)}\right)\right] \cdot \frac{k^2 g^4(k)}{X^4} \cdot \frac{g^4(X^2)}{g^4(k)} \frac{f(k)}{f\left(\frac{X^2}{g^2(X^2)}\right)} \\ &\quad \cdot I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq C \sum_{k=1}^{\infty} E\left[\frac{X^4}{g^4(X^2)} f\left(\frac{X^2}{g^2(X^2)}\right)\right] I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq CE\left[\frac{X^4}{g^4(X^2)} f\left(\frac{X^2}{g^2(X^2)}\right)\right] < \infty. \end{aligned}$$

Thus the proof of (2.1) is complete.

By the same argument as in (2.1), if (C3) and (C4) hold, we obtain

$$\begin{aligned} &\sum_{n=1}^{\infty} \frac{f(n)}{\sqrt{ng^3(n)}} E|X|^3 I\{|X| \leq \sqrt{ng(n)}\} \\ &\leq C \sum_{k=1}^{\infty} E|X|^3 I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \sum_{n=k}^{\infty} \frac{f(n)}{\sqrt{ng^3(n)}} \\ &\leq C \sum_{k=1}^{\infty} \frac{\sqrt{k}f(k)}{g^3(k)} E|X|^3 I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq C \sum_{k=1}^{\infty} \frac{kf(k)}{g^2(k)} E|X|^2 \\ &\quad \cdot I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq C \sum_{k=1}^{\infty} E|X|^2 I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \leq CE|X|^2 < \infty, \end{aligned}$$

if (C3) and (C4\*) hold, we obtain

$$\begin{aligned} &\sum_{n=1}^{\infty} \frac{f(n)}{\sqrt{ng^3(n)}} E|X|^3 I\{|X| \leq \sqrt{ng(n)}\} \\ &\leq C \sum_{k=1}^{\infty} \frac{\sqrt{k}f(k)}{g^3(k)} E|X|^3 I\{\sqrt{k}g(k) < |X| \leq \sqrt{k+1}g(k+1)\} \end{aligned}$$

$$\begin{aligned} &\leq C \sum_{k=1}^{\infty} E\left[\frac{X^4}{g^4(X^2)} f\left(\frac{X^2}{g^2(X^2)}\right)\right] \cdot \frac{\sqrt{kg(k)}}{|X|} \cdot \frac{g^4(X^2)}{g^4(k)} \frac{f(k)}{f\left(\frac{X^2}{g^2(X^2)}\right)} \\ &\quad \cdot I\{\sqrt{kg(k)} < |X| \leq \sqrt{k+1}g(k+1)\} \\ &\leq CE\left[\frac{X^4}{g^4(X^2)} f\left(\frac{X^2}{g^2(X^2)}\right)\right] < \infty. \end{aligned}$$

Thus the proof of (2.2) is complete.  $\square$

LEMMA 3. Under the conditions of Theorem 1, we have

$$\sum_{n=1}^{\infty} f(n)P\{|\mathcal{N}| > \frac{\beta g(n) \pm 1/g(n)}{\sigma_n}\} < \infty \text{ or } = \infty \tag{2.3}$$

according as

$$\sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2}{2}g^2(n)(1 + \alpha(n))\right\} < \infty \text{ or } = \infty \tag{2.4}$$

where  $\sigma_n^2 = E(XI\{|X| \leq \sqrt{ng(n)}\}) - EXI\{|X| \leq \sqrt{ng(n)}\})^2$ ,  $\alpha(n) = EX^2I\{|X| > \sqrt{ng(n)}\}/EX^2I\{|X| \leq \sqrt{ng(n)}\}$ ,  $\beta > 0$ .

*Proof.* It is easy to show that  $EX = 0$  and

$$\begin{aligned} &\left| \frac{g^2(n)}{\sigma_n^2} - \frac{g^2(n)}{EX^2I\{|X| \leq \sqrt{ng(n)}\}} \right| = \frac{g^2(n)|EXI\{|X| > \sqrt{ng(n)}\}|^2}{\sigma_n^2 EX^2I\{|X| \leq \sqrt{ng(n)}\}} \\ &\leq \frac{1}{\sigma_n^2 EX^2I\{|X| \leq \sqrt{ng(n)}\}} \frac{1}{n} \rightarrow 0. \end{aligned}$$

Note that  $P\{|\mathcal{N}| > x\} \sim \frac{2}{\sqrt{2\pi}} \frac{1}{x} \exp\{-\frac{x^2}{2}\}$  as  $x \rightarrow \infty$ ,  $\sigma_n \rightarrow 1$ ,  $1 = EX^2 = EX^2I\{|X| \leq \sqrt{ng(n)} + EX^2I\{|X| > \sqrt{ng(n)}\}$ , then we have

$$\begin{aligned} &\sum_{n=1}^{\infty} f(n)P\{|\mathcal{N}| > \frac{\beta g(n) \pm 1/g(n)}{\sigma_n}\} \\ &\sim C \sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2 g^2(n) \pm 2\beta + 1/g^2(n)}{2\sigma_n^2}\right\} \\ &\sim C \sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2 g^2(n)}{2\sigma_n^2}\right\} \\ &\sim C \sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2 g^2(n)}{2EX^2I\{|X| \leq \sqrt{ng(n)}\}}\right\} \\ &\sim C \sum_{n=1}^{\infty} \frac{f(n)}{g(n)} \exp\left\{-\frac{\beta^2 g^2(n)}{2}(1 + \alpha(n))\right\}, \end{aligned}$$



therefore the proof is completed. It is easy to see that  $\sum_{n=1}^{\infty} f(n)P\{|\mathcal{N}| > \frac{\beta g(n)+1/g(n)}{\sigma_n}\}$  and  $\sum_{n=1}^{\infty} f(n)P\{|\mathcal{N}| > \frac{\beta g(n)-1/g(n)}{\sigma_n}\}$  have the same convergence by the above proof.  $\square$

*Proof of Theorem 1.* Define  $S'_n = \sum_{k=1}^n X_k I\{|X_k| \leq \sqrt{ng}(n)\}$ , Obviously

$$\begin{aligned} |ES'_n| &= n|EXI\{|X| > \sqrt{ng}(n)\}| \leq n(EX^2)^{1/2}P^{1/2}\{|X| > \sqrt{ng}(n)\} \\ &\leq n\left(\frac{EX^2}{ng^2(n)}\right)^{1/2} = \frac{\sqrt{n}}{g(n)}. \end{aligned} \tag{2.5}$$

Note that

$$\begin{aligned} \{|S_n| > \beta\sqrt{ng}(n)\} &\subset \{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} \cup \{|S'_n| > \beta\sqrt{ng}(n)\}, \\ \{|S'_n| > \beta\sqrt{ng}(n)\} &\subset \{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} \cup \{|S_n| > \beta\sqrt{ng}(n)\}, \end{aligned}$$

and by (2.5)

$$\begin{aligned} \{|S'_n - ES'_n| > \beta\sqrt{ng}(n) + \frac{\sqrt{n}}{g(n)}\} &\subset \{|S'_n| > \beta\sqrt{ng}(n)\} \\ &\subset \{|S'_n - ES'_n| > \beta\sqrt{ng}(n) - \frac{\sqrt{n}}{g(n)}\}. \end{aligned}$$

Hence

$$\begin{aligned} &P\{|S'_n - ES'_n| > \beta\sqrt{ng}(n) + \frac{\sqrt{n}}{g(n)}\} - P\{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} \leq P\{|S_n| > \beta\sqrt{ng}(n)\} \\ &\leq P\{|S'_n - ES'_n| > \beta\sqrt{ng}(n) - \frac{\sqrt{n}}{g(n)}\} + P\{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\}. \end{aligned}$$

By the fact  $|a| - |a - b| \leq |b| \leq |a| + |a - b|$ , then

$$\begin{aligned} &P\{|\mathcal{N}| > \frac{\beta\sqrt{ng}(n) + \frac{\sqrt{n}}{g(n)}}{\sqrt{\text{Var}(S'_n)}}\} - |P\{|S'_n - ES'_n| > \beta\sqrt{ng}(n) + \frac{\sqrt{n}}{g(n)}\} \\ &\quad - P\{|\mathcal{N}| > \frac{\beta\sqrt{ng}(n) + \frac{\sqrt{n}}{g(n)}}{\sqrt{\text{Var}(S'_n)}}\}| - P\{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} \\ &\leq P\{|S_n| > \beta\sqrt{ng}(n)\} \\ &\leq P\{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} + |P\{|S'_n - ES'_n| > \beta\sqrt{ng}(n) - \frac{\sqrt{n}}{g(n)}\} \\ &\quad - P\{|\mathcal{N}| > \frac{\beta\sqrt{ng}(n) - \frac{\sqrt{n}}{g(n)}}{\sqrt{\text{Var}(S'_n)}}\}| + P\{|\mathcal{N}| > \frac{\beta\sqrt{ng}(n) - \frac{\sqrt{n}}{g(n)}}{\sqrt{\text{Var}(S'_n)}}\}. \end{aligned}$$

By (2.1), we have

$$\sum_{n=1}^{\infty} f(n)P\{\max_{1 \leq k \leq n} |X_k| > \sqrt{ng}(n)\} \leq \sum_{n=1}^{\infty} nf(n)P\{|X| > \sqrt{ng}(n)\} < \infty.$$

By Lemma 1 and (2.2),  $\text{Var}(S'_n) = n\sigma_n^2$ ,  $\sigma_n \rightarrow 1$ ,  $\frac{\beta g(n) \pm \frac{1}{g(n)}}{\sigma_n} \rightarrow \beta g(n)$ , we know

$$\begin{aligned} & \sum_{n=1}^{\infty} f(n) |P\{|S'_n - ES'_n| > \beta \sqrt{n}g(n) \pm \frac{\sqrt{n}}{g(n)}\} - P\{|\mathcal{N}| > \frac{\beta \sqrt{n}g(n) \pm \frac{\sqrt{n}}{g(n)}}{\sqrt{\text{Var}(S'_n)}}\}| \\ &= \sum_{n=1}^{\infty} f(n) |P\{|S'_n - ES'_n| > \beta \sqrt{n}g(n) \pm \frac{\sqrt{n}}{g(n)}\} - P\{|\mathcal{N}| > \frac{\beta g(n) \pm \frac{1}{g(n)}}{\sigma_n}\}| \\ &\leq C \sum_{n=1}^{\infty} \frac{f(n)}{\sqrt{n}g^3(n)} E|X|^3 I\{|X| \leq \sqrt{n}g(n)\} < \infty. \end{aligned}$$

Therefore

$$\sum_{n=1}^{\infty} f(n) P\{|S_n| > \beta \sqrt{n}g(n)\} < \infty \text{ or } = \infty \quad (2.6)$$

according as

$$\sum_{n=1}^{\infty} f(n) P\{|\mathcal{N}| > \frac{\beta g(n) \pm \frac{1}{g(n)}}{\sigma_n}\} < \infty \text{ or } = \infty, \quad (2.7)$$

then by Lemma 3, the proof of Theorem 1 is complete.  $\square$

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*Yong Zhang*  
*School of Mathematics*  
*Jilin University*  
*Changchun 130012, P.R.China*  
*e-mail: zyong2661@jlu.edu.cn*